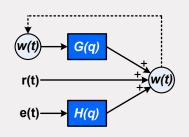
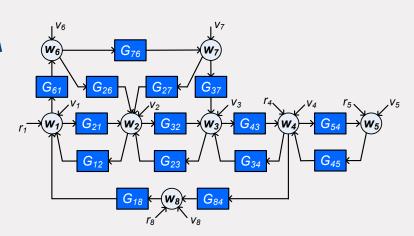


From classical models...

Nodes: input and output



...to dynamic network models





$$w(t) = G(q)w(t) + R(q)r(t) + H(q)e(t)$$

**Derivation of the predictor** - In case of full rank disturbances (H(q) square and invertible)\*:

$$e(t) = H(q)^{-1}[(I - G(q))w(t) - R(q)r(t)]$$

$$\begin{array}{lll} w(t) & = & G(q)w(t) + R(q)r(t) + (H(q) - I) \textcolor{red}{e(t)} + e(t) \\ & = & G(q)w(t) + R(q)r(t) + (I - H(q)^{-1})[(I - G(q))w(t) - R(q)r(t)] + e(t) \\ & = & [I - H(q)^{-1}(I - G(q))]w(t) + [H(q)^{-1}R(q)]r(t) + e(t) \end{array}$$

$$\hat{w}(t|t-1) = \bar{E}\{w(t) \mid w^{t-1}, r^t\} = [I - H(q)^{-1}(I - G(q))]w(t) + [H(q)^{-1}R(q)]r(t)$$
 
$$= [\underbrace{I - H(q)^{-1}}_{"output"} + \underbrace{H(q)^{-1}G(q)}_{"input"}]w(t) + H(q)^{-1}R(q)r(t)$$



 $<sup>^</sup>st$  For simplicity we assume G strictly proper

$$w(t) = G(q)w(t) + R(q)r(t) + H(q)e(t)$$

Predictor model:

$$\hat{w}(t|t-1;\theta) = [I-H(q,\theta)^{-1} + H(q,\theta)^{-1}G(q,\theta)]w(t) + H(q,\theta)^{-1}R(q,\theta)r(t)$$

This leads to a prediction error:

$$arepsilon(t, heta) = H(q, heta)^{-1}[(I-G(q, heta))w(t)-R(q, heta)r(t)]$$

and a prediction error estimator:

$$\hat{ heta}_N = rg \min_{ heta} rac{1}{N} \sum_{t=0}^{N-1} arepsilon(t, heta)^T Q arepsilon(t, heta)$$
  $Q>0$ 



This is a consistent estimator, under the following conditions:

- System is in the model set,  $\mathcal{S} \in \mathcal{M}$
- Model set  $\mathcal M$  is globally network identifiable at  $\mathcal S$
- There are no algebraic loops in the network (every loop has a delay)
- The present r signals are persistently exciting of a sufficiently high order (data-informativity)

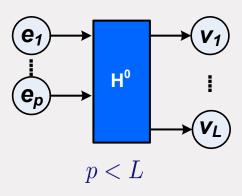
If disturbances are uncorrelated, i.e.  $H(q,\theta)$  diagonal: problem decomposed in L MISO problems



What can we do if the disturbances are not full rank?

i.e.  $\Phi_v(\omega)$  does not have full rank for all  $\omega$ 

In large scale networks there may be common sources behind multiple disturbances



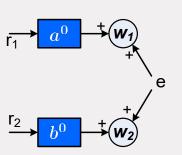
This situation is not really treated in the classical PEM literature

but known in (dynamic) factor analysis<sup>[1]</sup>



A simple example: 2 nodes disturbed by 1 noise

$$\varepsilon_1(t, a) = w_1(t) - ar_1(t)$$
  
$$\varepsilon_2(t, b) = w_2(t) - br_2(t)$$



Estimate  $a^0$  with parameter a by minimizing  $\frac{1}{N}\sum_{t=0}^{N-1}\varepsilon_1(t,\theta)^2$ 

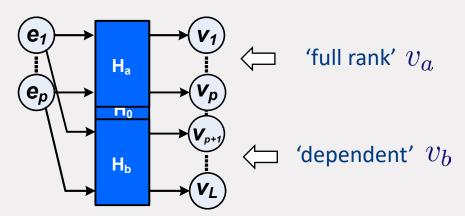
but note that using the dependency  $\varepsilon_1(t,a^0)=\varepsilon_2(t,b^0)$  for the model leads to

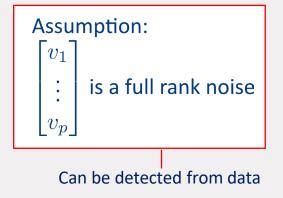
$$\varepsilon_1(t,a) = \varepsilon_2(t,b) \Leftrightarrow e + (a^0 - a)r_1 = e + (b^0 - b)r_2$$

which for persistently excitating and independent r signals gives variance-free estimates  $\hat{a}=a^0$  and  $\hat{b}=b^0$ .



Noise can have  $\dim(e) \leq \dim(v)$ 





$$H=egin{bmatrix} H_a \ H_b \end{bmatrix}$$
 Ha is monic, i.e.  $\lim_{z o\infty}H_a(z)=I$  Hb is non-square



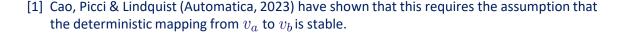
Particular spectral factorization result of a stochastic process v with rank  $p \leq L$  and with  $(v_1 \cdots v_p)$  full rank:

$$\Phi_v = \breve{H} \breve{\Delta} \breve{H}^*$$

$$\text{with } \breve{H} = \begin{bmatrix} H_a & 0 \\ H_b - \Gamma & I \end{bmatrix}, \ \breve{\Delta} = \begin{bmatrix} I \\ \Gamma \end{bmatrix} \Delta \begin{bmatrix} I \\ \Gamma \end{bmatrix}^T, \ \Delta > 0, \quad \Gamma = \lim_{z \to \infty} H_b(z)$$

Consequently: 
$$v(t) = \breve{H}^0(q) e(t) = \begin{bmatrix} H_a^0 & 0 \\ H_b^0 - \Gamma^0 & I \end{bmatrix} \begin{bmatrix} e \\ \Gamma^0 e \end{bmatrix}$$

with 
$$\begin{vmatrix} e_a \\ e_b \end{vmatrix} := \breve{H}^0(q)^{-1}v(t)$$
 it follows that  $\Gamma^0e_a(t) - e_b(t) = 0$  for all  $t$ 





$$\varepsilon(t,\theta) = \begin{bmatrix} \varepsilon_a(t,\theta) \\ \varepsilon_b(t,\theta) \end{bmatrix} = \begin{bmatrix} H_a(q,\theta) & 0 \\ H_b(q,\theta) - \Gamma(\theta) & I \end{bmatrix}^{-1} [(I - G(q,\theta))w(t) - R(q,\theta)r(t)]$$

with the constraint:  $\Gamma(\theta)\varepsilon_a(t,\theta)-\varepsilon_b(t,\theta)=0$  for all t

Weighted least-squares method: discard dependencies in noise:

$$\theta^{\star} = \arg\min_{\theta} \mathbb{E} \ \varepsilon^{T}(t, \theta) \ Q \ \varepsilon(t, \theta) \qquad \qquad Q > 0$$

leads to consistent estimates of the network, under the same conditions as for the full rank case

**However:** For ML/ minimum variance results we typically would need  $Q=[cov(e)]^{-1}$  but cov(e) is not invertible



**Constrained least-squares method**: include the noise constraint:

$$\begin{array}{ll} \theta^{\star} = \arg\min_{\theta} \bar{\mathbb{E}} \; \{ \varepsilon_a^T(t,\theta) \; Q_a \; \varepsilon_a(t,\theta) \} & Q_a > 0 \\ \\ \text{subject to } \bar{\mathbb{E}} \{ Z^T(t,\theta) Z(t,\theta) \} = 0 \\ \\ \text{with } Z(t,\theta) := \Gamma(\theta) \varepsilon_a(t,\theta) - \varepsilon_b(t,\theta) \end{array}$$

leads to consistent estimates of the network, under the same conditions as before

This approach provides Maximum Likelihood Estimates, through constrained optimization of the log-likelihood function



Constraint can be infeasible  $\implies$  Relax the constraint

$$\begin{split} \theta^\star &= \arg\min_{\theta} \bar{\mathbb{E}} \; \{ \varepsilon_a^T(t,\theta) \; Q_a \; \varepsilon_a(t,\theta) \} + \lambda \bar{\mathbb{E}} \{ Z^T(t,\theta) Z(t,\theta) \} \\ \text{with } Z(t,\theta) &:= \Gamma(\theta) \varepsilon_a(t,\theta) - \varepsilon_b(t,\theta) \\ \text{and } \lambda \text{ a tuning parameter} \end{split}$$

Which is a WLS with parameterized weight  $\theta^\star = \arg\min_{\theta} \bar{\mathbb{E}} \left\{ \varepsilon_a^T(t,\theta) \; Q_\lambda \; \varepsilon_a(t,\theta) \right\}$ 

$$Q_{\lambda}(\Gamma) = \begin{bmatrix} Q_a + \lambda \Gamma^T(\theta) \Gamma(\theta) & -\lambda \Gamma^T(\theta) \\ \lambda \Gamma(\theta) & \lambda I \end{bmatrix}$$

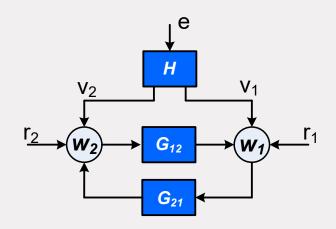
Computationally more attractive



# **Network identification – simulation example**

$$N=1000$$
 samples, 100 realizations of data,  $\sigma_e^2=100\sigma_{r_1}^2=100\sigma_{r_2}^2$ 

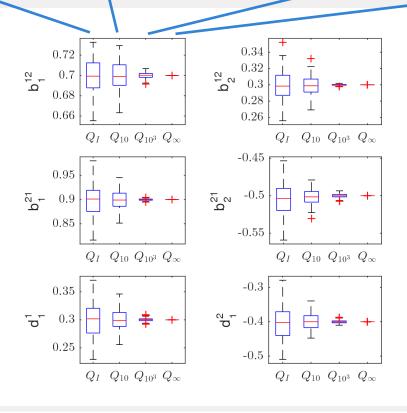
$$G_{ij}(q,\theta_0) = b_1^{ij}q^{-1} + b_2^{ij}q^{-2}$$
$$H_i(q,\theta_0) = \frac{1}{1 + d_1^i q^{-1}}$$



6 parameters in total



WLS: Q = I, Relaxed  $\lambda = 10$ , Relaxed  $\lambda = 1000$ , CLS



Variance-free when dependencies taken into account



**However:** methods based on non-convex optimization scale poorly to larger dimensions

**Alternatives:** Sequential / multi-step methods / weighted nullspace fitting<sup>[1]-[3]</sup>

- Estimate a (regularized) high-order ARX model
- Either reconstruct the innovation signal to become a measured input, or approximate the high-order model with linear techniques

[3] Fonken et al, Automatica, 2022

Iterate to find the optimal criterion weighting for optimal variance

To be further developed to arrive at robust algorithms



<sup>[1]</sup> Galrinho, Rojas and Hjalmarsson, TAC 2019

<sup>[2]</sup> Weerts et al, SYSID 2018